



## Performance of Seven Statistics for Mean Difference Testing Between Two Populations Under Combined Assumption Violations

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### Abstract

The objective of this research was to compare the performance of seven statistics for mean difference testing between two populations when data did not follow assumptions whereas the simulation conditions were determined as 5 distributions, variance, and sample size in both cases are equal and unequal. The results showed that when the population had log-normal distribution, gamma distribution and poisson distribution and equal variance, the Welch Based on Rank test (WBR test) were most effective. When the population had log-normal distribution, gamma distribution, exponential distribution, poisson distribution and uniform distribution and unequal variance, the Welch t test was distinctively found to have a higher performance than others testing statistics.

**Keywords:** Two-sample location test, Parametric test, Non-parametric test, Non-parametric bootstrap test, t test, Welch t test, Welch Based on Rank test, Brunner-Munzel test, Yuen-Welch test, Exact Wilcoxon signed-rank test

### Introduction

The testing of differences on location or central tendency values between two independent populations is widely used for hypothesis testing in several research fields, such as educational research, medicine, business administration, etc. The said hypothesis testing normally applies an independent t test when both populations have a normal distribution and equal variances. However, if both populations have a normal distribution and unequal variances, the Welch t test is used. These statistics are in the parametric test (Stonehouse & Forrester, 1998; Fagerland & Sandvik, 2009; Ahad, Othman, & Yahaya, 2011). However, if the populations have a non-normal distribution, the hypothesis test will be performed using nonparametric statistics. The most popular testing statistic is Wilcoxon-Mann-Whitney (Mann, & Whitney, 1947; Harwell & Serlin, 1989; Fagerland & Sandvik, 2009). It could be said that the said testing statistics, in many conditions, experience problems with efficiency, especially when the populations have an unequal variance (Harwell, Rubinstein, Hayes, & Olds, 1992; Zimmerman & Zumbo, 1993a, 1993b; Stonehouse & Forrester, 1998). Afterwards, the testing statistics are adjusted to what is called Brunner-Munzel (BM). It aims to present the distinction when there are unequal variances, known as the Behrences-Fisher problem (Brunner & Munzel, 2000). Nevertheless, there is an argument about the effectiveness of the above testing statistics (Medina, Kimberg, Chatterjee, & Coslett, 2010). Also, the Welch t test is used for hypothesis testing by transforming data into ranking order (Welch, 1938; Reiczigel, Zakarias, & Rozsa, 2005; Winter, 2013).

In addition, testing statistics are developed to test the mean difference when the population experiences a skewed distribution or outlier value by using the trimmed means. In detail, before finding the mean, the data is in ascending order and the smallest and highest values are cut out. It could be said that the maximum efficiency is found by cutting out 20 percent of the data (Wilcox, 1994; Wicox, 2005). This comparison of



trimmed means is called the Yuen-welch test statistics (Yuen, 1974). The study of Fagerland and Sandvik (2009) revealed that the above test statistics are suitable for analysis with a positive skewed distribution. For hypothesis testing relating the mean difference, the standard error could be adjusted by using the non-parametric bootstrap method, in which the algorithm was proposed by Efron & Tibshirani (1993). The outstanding feature of this method is that there is no assumption on the normal distribution and the equality of variance because it is done by replacing the sampling for estimating the parameters.

In many research fields, data sometimes may not meet the assumptions on normal distribution and they mostly occur with unequal variances (Wilcox, 1990; Wilcox & Keselman, 2003). For instance, for research in medical fields, data are normally skewed (Bridge & Sawilowsky, 1999). If two skewed distribution have unequal locations, it can be expected that the variances will also differ. Later on, medical data regularly show an integration of skewness and unequal variances (Fagerland & Sandvik, 2009). Therefore, data analysis using parametric statistics is not proper. The objective of this research is to compare the efficiency of the test statistic both parametric statistics and non-parametric statistics, such as the Welch Based on Rank test (WBR test), Yuen-Welch test (Y-W test), Brunner-Munzel test (B-M test) and Exact Wilcoxon signed-rank test. The test statistic is done by considering the efficiency of capacity to control type I errors and testing power. The capacity to control type I errors is based on Bradley (1978) criteria by testing the two-tail hypothesis at a significant level of 0.05 and 0.01. If the values range between 0.025–0.075 or 0.005–0.015 respectively, it is considered to have the capacity to control a type I error.

## Methods and Materials

### Statistical Method

The test statistic for the hypothesis test on different locations of two independent populations were as follows:

#### 1. Independent sample t test (t test)

The independent sample t test (t test) was used for parametric statistics with assumptions consisting of normal distribution and equal variances population. The details are: (Fagerland & Sandvik, 2009)

$$t = \frac{\bar{X}_1 - \bar{X}_2}{\sqrt{\frac{S_p^2}{n_1} + \frac{S_p^2}{n_2}}} \quad (1)$$

Where  $t$  is t-distribution

$$S_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}$$

The reject  $H_0$  is  $t \leq -t_{\alpha/2, n_1+n_2-1}$  or  $t \geq t_{\alpha/2, n_1+n_2-1}$

#### 2. Welch t test

The Welch t test was used for parametric statistics with the assumption of normal distribution. The Welch t test was proposed by Welch (1937) for unequal variances. The details are:



$$\text{Welch} = \frac{\bar{X}_1 - \bar{X}_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}} \quad (2)$$

Where Welch is t-distribution

$$S_1^2 = \frac{\sum_{i=1}^{n_1} (X_{i1} - \bar{X}_1)^2}{n_1 - 1}$$

$$S_2^2 = \frac{\sum_{i=1}^{n_2} (X_{i2} - \bar{X}_2)^2}{n_2 - 1}$$

$$\text{The reject } H_0 \text{ is Welch} \leq -t_{\alpha/2, df} \text{ or Welch} \geq t_{\alpha/2, df} \text{ by } df = \frac{\left( \frac{S_1^2}{n_1} + \frac{S_2^2}{n_2} \right)^2}{\left( \frac{S_1^2}{n_1} \right)^2 + \left( \frac{S_2^2}{n_2} \right)^2}$$

$$n_1 - 1 \quad n_2 - 1$$

### 3. t-test based on bootstrap standard errors (t-BSE test)

Hypothesis testing with t-BSE was described as follows (Efron & Tibshirani, 1993):

- 1) Evaluated  $\hat{\theta} = \bar{X}_1 - \bar{X}_2$
- 2) Evaluated the standard error by replacement sampling at sample sizes  $n_1^*$  and  $n_2^*$  from  $n_1 + n_2$  and evaluated  $\hat{\theta}^*(b) = \bar{X}_1^* - \bar{X}_2^*$
- 3) Evaluated  $\hat{se}_B = \left\{ \sum_{b=1}^B [\hat{\theta}^*(b) - \hat{\theta}^*(.)]^2 / B - 1 \right\}^{1/2}$ , for this paper, Bootstrap was 1,000 time,
- where  $\hat{\theta}^*(.) = \frac{\sum_{b=1}^B \hat{\theta}^*(b)}{B}$
- 4) Evaluated  $t - BSE = \frac{\hat{\theta}}{\hat{se}_B}$  (3)
- 5) Rejected  $H_0$  when  $t - BSE \leq -t_{\alpha/2, B-1}$  or  $t - BSE \geq t_{\alpha/2, B-1}$

### 4. Welch Based on Rank test (WBR test)

Hypothesis testing with Welch Based on Rank was described as follows: ( Welch, 1938; Reiczigel, et al., 2005; Winter, 2013).

- 1) Combined  $n_1$  and  $n_2$  and ascended the data, if tied values was found, the midranks were averaged.
- 2) Calculated the mean and standard deviation of midrank in the first and second samples as  $\bar{R}_1$  and  $\bar{R}_2$ , and calculated  $S_1^2$  and  $S_2^2$  respectively

$$3) \text{Evaluated the test statistic } WBR = \frac{\bar{R}_1 - \bar{R}_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}} \quad (4)$$



4) Rejected  $H_0$  when  $WBR \leq -t_{\alpha/2, df}$  or  $WBR \geq t_{\alpha/2, df}$  by  $df = \frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}\right)^2}{\frac{(S_1^2/n_1)^2}{n_1-1} + \frac{(S_2^2/n_2)^2}{n_2-1}}$

##### 5. Brunner–Munzel test (B–M test)

Brunner and Munzel (2000) proposed the hypothesis testing method to solve the Behrens–Fisher problem or unequal variance. Following the notation  $R_1$  and  $R_2$  are the midranks of  $n_1$  and  $n_2$  based on pooling all observations,  $V_1$  and  $V_2$  were the midranks within each sample.  $\bar{R}_1$  and  $\bar{R}_2$  were the means of the pooled midranks. The test statistics were detailed as follows:

$$B-M = \frac{\bar{R}_1 - \bar{R}_2}{(n_1 + n_2) \sqrt{\frac{SB_1^2}{n_1 n_2} + \frac{SB_2^2}{n_1^2 n_2}}} \quad (5)$$

where

$$SB_1^2 = \frac{1}{n_1 - 1} \sum_{i=1}^{n_1} \left( R_1^i - V_1^i - \bar{R}_1 + \frac{n_1 + 1}{2} \right)^2$$

$$SB_2^2 = \frac{1}{n_2 - 1} \sum_{i=1}^{n_2} \left( R_2^i - V_2^i - \bar{R}_2 + \frac{n_2 + 1}{2} \right)^2$$

Reject  $H_0$  when  $B-M \leq -t_{\alpha/2, df}$  or  $B-M \geq t_{\alpha/2, df}$  by

$$df = \frac{\left( \frac{SB_1^2}{n_1} + \frac{SB_2^2}{n_2} \right)^2}{\frac{SB_1^4}{n_1^2(n_2 - 1)} + \frac{SB_2^4}{n_2^2(n_1 - 1)}}$$

##### 6. Yuen–Welch test (Y–W test)

Yuen–Welch test used trimmed means by cutting of 20% or  $\gamma = 0.2$ , which is the best option (Wilcox, 1994; Wilcox, 2005). Specifically, the least and greatest values were reduced by 20% and defined  $g_1 = \gamma n_1$  and  $g_2 = \gamma n_2$  as trimmed observations from each tail in  $n_1$  and  $n_2$ . Indicated the number of remaining observations in the trimmed samples by  $h_1 = n_1 - 2g_1$  and  $h_2 = n_2 - 2g_2$ .  $SW_1^2$  and  $SW_2^2$  were sample winsorized variances in each sample. This test statistic could be explained as follows (Yuen, 1974):

$$Y-W = \frac{\bar{X}_{\gamma 1} - \bar{X}_{\gamma 2}}{\sqrt{d_1 + d_2}} \quad (6)$$

where

$$d_1 = \frac{SW_1^2(n_1 - 1)}{h_1(h_1 - 1)}$$

$$d_2 = \frac{SW_2^2(n_2 - 1)}{h_2(h_2 - 1)}$$

$$SW_1^2 = \frac{1}{n_1 - 1} \sum_{i=1}^{n_1} (W_1^i - \bar{W}_1)^2$$

$$SW_2^2 = \frac{1}{n_2 - 1} \sum_{i=1}^{n_2} (W_2^i - \bar{W}_2)^2$$



Rejected  $H_0$  when  $Y-W \leq -t_{\alpha/2, df}$  or  $Y-W \geq t_{\alpha/2, df}$  by

$$df = \frac{(d_1 + d_2)^2}{\left( \frac{d_1^2}{h_1-1} + \frac{d_2^2}{h_2-1} \right)}$$

### 7) Exact Wilcoxon signed-rank test

The Exact Wilcoxon signed-rank test was proposed by Wilcoxon (1945). The Exact Wilcoxon signed-rank test can apply the exact permutation distribution of ranks since the exact test is only practicable for small sample. However, due to the current advances in computers, P-value is calculated from the actual distribution except when there are duplicates (tied) and approximated by normal distribution. For computation of such test statistics, WRS2 in R program command is used.

The Exact Wilcoxon signed-rank test were designed based on W rank statistics used for ordering the samples which are the combination of  $N = n_1 + n_2$   $X_1$ -value and  $X_2$ -value ranking from least to highest determining  $S_1$  as the rank of  $X_{21}, \dots, X_{2n_2}$  as the rank of  $X_{2n_2}$  in this joint ordering. That is, W refers to the sum of the ranks determined for the  $X_2$ -value.

$$W = \sum_{j=1}^{n_2} S_j \quad (7)$$

Reject  $H_0$  when  $W \leq n_2(n_1 + n_2 + 1)/2 - W_{\alpha/2}$  or  $W \geq W_{\alpha/2}$

## Methodology

The Monte Carlo simulation was used to estimate the type I error and the power of the test. R program Version 3.5.2 was written to generate data for the following conditions. In each case, 10,000 iterations of data sets were generated.

### 1. Generating of two population data groups into five distributions.

#### 1) Population with log-normal distribution

$$X_1 \sim \text{log Norm}(\mu = 0, \sigma = 1) - 1.65 \quad (8)$$

$$X_2 \sim \sigma_2(\text{log Norm}(\mu = 0, \sigma = 1) - 1.65 - \text{shift}) \quad (9)$$

#### 2) Population with gamma distribution

$$X_1 \sim \text{gamma}(\alpha = 1, \beta = 1) - 1 \quad (10)$$

$$X_2 \sim \sigma_2(\text{gamma}(\alpha = 1, \beta = 1) - 1 - \text{shift}) \quad (11)$$

#### 3) Population with exponential distribution

$$X_1 \sim \text{exp}(\lambda = 0.5) - 0.5 \quad (12)$$

$$X_2 \sim \sigma_2(\text{exp}(\lambda = 0.5) - 0.5 - \text{shift}) \quad (13)$$

#### 4) Population with poission distribution

$$X_1 \sim \text{pois}(\lambda = 5) - 5 \quad (14)$$

$$X_2 \sim \sigma_2(\text{pois}(\lambda = 5) - 5 - \text{shift}) \quad (15)$$

#### 5) Population with uniform distribution

$$X_1 \sim \text{unif}(\min = 0, \max = 1) - 0.5 \quad (16)$$



$$X_2 \sim \sigma_2(\text{unif}(\min = 0, \max = 1) - 0.5 - \text{shift}) \quad (17)$$

When determine shift = 0 as the calculation for type I error and shift = 1, 1, 0.5, 1 and 0.25 when population had log-normal distribution, gamma distribution, exponential distribution, poission distributionand uniform distribution, respectively, was the calculation of power test.

2. Determining four variance ratios for both equal and unequal variance, which were 1:1, 1:2.25, 1:4, and 1:9.

3. Determiningthe sample sizes for both equal and unequal sizes consisting of seven sizes, which were (10,10), (10,30), (30,10), (30,30), (50,100), (100,50), and (100,100).

4. Finding the probability of the type I error by assigning the equivalent mean ratio and counting the rejection number of  $H_0$ , then dividing it by the 10,000 iteration rounds for each condition. This hypothesis testing was done at a significant level of 0.05 and 0.01. If the probability of a type I error was between 0.025–0.075 or 0.005–0.015 respectively, it would have been considered to have the ability to control the type I error based on criteria from Bradley (1978).

5. Finding the testing power by assigning the difference of the mean based on the effect size. If it is above 0.80, it is considered to have a high effect (Cohen, 1988). The testing power was calculated by counting the rejection number of  $H_0$  and dividing it by 10,000. This hypothesis testing was done at a significant level of 0.05 and 0.01.

6. Analyzing the performance of each situation test statistics that can control type I error and has the highest Power of the test

## Results

The results of simulation were reported in following order in type I errors and power of the test for testing statistics had the following details:

**Table 1** The probability of a type I error and power test (0.05 significant level) of statistics on location testing between two populations with log-normal distribution classified by sample size and variance

Sample Size	Var.ratio	t		Welch		t-BSE		WBR		B-M		Y-W		Wilcoxon	
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT
10:10	1:1	<b>0.0851</b>	0.3092	<b>0.0274</b>	0.2960	<b>0.0383</b>	0.3174	<b>0.0551</b>	<b>0.5339</b>	<b>0.0579</b>	0.4983	<b>0.0302</b>	0.3860	<b>0.0438</b>	0.5026
	1:2.25	<b>0.0517</b>	0.4304	<b>0.0460</b>	0.4200	<b>0.0559</b>	0.4387	0.1174	0.6589	0.1103	0.5809	<b>0.0692</b>	<b>0.5873</b>	0.1040	0.6361
	1:4	0.0855	0.5037	0.0802	0.4908	0.0889	0.5093	0.1932	0.7127	0.1676	0.6106	0.1400	0.6713	0.1755	0.6969
	1:9	0.1292	0.5460	0.1221	0.5308	0.1328	0.5514	0.2552	0.7385	0.1982	0.5976	0.2202	0.7132	0.2404	0.7271
10:30	1:1	<b>0.0409</b>	0.3634	<b>0.0697</b>	0.2783	<b>0.0488</b>	0.3661	<b>0.0507</b>	<b>0.8950</b>	<b>0.0549</b>	0.8815	<b>0.0585</b>	0.5362	<b>0.0457</b>	0.7738
	1:2.25	<b>0.0364</b>	0.4207	<b>0.0382</b>	0.5105	<b>0.0370</b>	0.4256	0.2027	0.9809	0.1965	0.9660	<b>0.0528</b>	<b>0.8814</b>	0.1109	0.9061
	1:4	<b>0.0561</b>	0.4417	<b>0.0376</b>	<b>0.6101</b>	<b>0.0363</b>	0.4448	0.3985	0.9880	0.3558	0.9739	0.1416	0.9461	0.2015	0.9314
	1:9	<b>0.0815</b>	0.4575	<b>0.0572</b>	<b>0.6785</b>	<b>0.0389</b>	0.4599	0.5778	0.9941	0.4832	0.9844	0.3003	0.9789	0.3071	0.9538
30:10	1:1	<b>0.0439</b>	0.3784	<b>0.0687</b>	0.4685	<b>0.0452</b>	0.3820	<b>0.0555</b>	0.5704	<b>0.0570</b>	0.5517	<b>0.0551</b>	0.5456	<b>0.0505</b>	<b>0.7008</b>
	1:2.25	0.0920	0.5714	0.1150	0.5149	0.0946	0.5747	0.1264	0.6607	0.1227	0.6183	0.1355	0.6406	0.2029	0.8084
	1:4	0.1515	0.6497	0.1282	0.5400	0.1550	0.6512	0.1811	0.6947	0.1659	0.6411	0.1814	0.6794	0.2987	0.8351
	1:9	0.2395	0.7011	0.1410	0.5435	0.2444	0.7027	0.2363	0.7227	0.2059	0.6404	0.2301	0.7125	0.3839	0.8556
30:30	1:1	<b>0.0379</b>	0.5474	<b>0.0365</b>	0.5453	<b>0.0393</b>	0.5505	<b>0.0468</b>	<b>9511.0</b>	<b>0.0480</b>	0.9411	<b>0.0384</b>	0.8395	<b>0.0468</b>	<b>0.9611</b>
	1:2.25	<b>0.0564</b>	0.6599	<b>0.0537</b>	0.6566	<b>0.0564</b>	<b>0.8815</b>	0.2799	0.9868	0.2640	0.9782	0.1293	0.9435	0.2799	0.9868
	1:4	<b>0.0690</b>	0.6965	<b>0.0676</b>	0.6925	<b>0.0708</b>	<b>0.8980</b>	0.4307	0.9909	0.3852	0.9848	0.2192	0.9685	0.4310	0.9910
	1:9	0.0953	0.7136	0.0939	0.7083	0.0972	0.7154	0.5832	0.9929	0.5078	0.9866	0.3558	0.9801	0.5848	0.9932
50:100	1:1	<b>0.0448</b>	0.7897	<b>0.0619</b>	0.8279	<b>0.0482</b>	0.7908	<b>0.0498</b>	<b>1.0000</b>	<b>0.0503</b>	<b>1.0000</b>	<b>0.0488</b>	0.9982	<b>0.0476</b>	<b>1.0000</b>
	1:2.25	<b>0.0288</b>	0.8496	<b>0.0484</b>	<b>0.9144</b>	<b>0.0280</b>	0.8491	0.5947	1.0000	0.5705	1.0000	0.1797	1.0000	0.4967	1.0000
	1:4	<b>0.0307</b>	0.8539	<b>0.0479</b>	<b>0.9247</b>	<b>0.0326</b>	0.8547	0.8878	1.0000	0.8554	1.0000	0.4562	1.0000	0.8120	1.0000
	1:9	<b>0.0303</b>	0.8779	<b>0.0618</b>	<b>0.9460</b>	<b>0.0312</b>	0.8788	0.9714	1.0000	0.9509	1.0000	0.7324	1.0000	0.9330	1.0000
100:50	1:1	<b>0.0487</b>	0.7920	<b>0.0446</b>	0.7704	<b>0.0467</b>	0.7925	<b>0.0494</b>	0.9978	<b>0.0600</b>	0.9975	<b>0.0470</b>	0.9793	<b>0.0497</b>	<b>0.9990</b>
	1:2.25	0.0820	0.8589	<b>0.0656</b>	<b>0.7999</b>	0.0832	0.8581	0.4281	0.9999	0.4116	0.9998	0.1859	0.9966	0.5049	1.0000
	1:4	0.1190	0.8747	0.0802	0.8049	0.1195	0.8751	0.6550	1.0000	0.6257	1.0000	0.3346	0.9991	0.7406	1.0000
	1:9	0.1653	0.9062	0.0970	0.8351	0.1658	0.9063	0.7962	1.0000	0.7578	1.0000	0.5166	0.9994	0.8705	1.0000

**Table 1** (Cont.)

Sample Size	Var.ratio	t		Welch		t-BSE		WBR		B-M		Y-W		Wilcoxon	
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT
100:100	1:1	<b>0.0489</b>	0.9055	<b>0.0485</b>	0.9051	<b>0.0492</b>	0.9051	<b>0.0541</b>	<b>1.0000</b>	<b>0.0548</b>	<b>1.0000</b>	<b>0.0493</b>	0.9999	<b>0.0540</b>	<b>1.0000</b>
	1:2.25	<b>0.0568</b>	0.9462	<b>0.0661</b>	0.9457	<b>0.0558</b>	<b>0.9463</b>	0.6843	1.0000	0.6601	1.0000	0.2612	1.0000	0.6836	1.0000
	1:4	<b>0.0860</b>	<b>0.9865</b>	<b>0.0646</b>	0.9355	<b>0.0860</b>	0.9361	0.9042	1.0000	0.8811	1.0000	0.5095	1.0000	0.9042	1.0000
	1:9	<b>0.0737</b>	<b>0.9883</b>	<b>0.0724</b>	0.9372	<b>0.0732</b>	0.9378	0.9718	1.0000	0.9549	1.0000	0.7361	1.0000	0.9718	1.0000

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 1, it could be found that when the populations had log-normal distribution and equal variance at 0.05 significant level, all test statistics could control type I errors in all sample sizes. For unequal variance cases Welch t test, t test and t-BSE test could control type I errors in almost all sample sizes, unless small sample size and higher variance another group.

It could be found that when the population had log-normal distribution and equal variance, the test statistics with the highest testing power that can control type I errors are WBR test (71.43%) and Exact Wilcoxon signed-rank test (71.43%). For unequal variance, the test statistics with the highest testing power that could control were Welch t test (28.57%), t-BSE test (14.29%) and t test (9.52%) respectively.

**Table 2** The probability of a type I error and power test (0.01 significant level) of statistics on location testing between two populations with log-normal distribution classified by sample size and variance

Sample Size	Var.ratio	t		Welch		t-BSE		WBR		B-M		Y-W		Wilcoxon	
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT
10:10	1:1	<b>0.0081</b>	0.1395	0.0031	0.1214	0.0038	0.1222	<b>0.0108</b>	<b>0.3081</b>	0.0159	0.3099	0.0047	0.1572	<b>0.0078</b>	0.2740
	1:2.25	<b>0.0101</b>	<b>0.2864</b>	<b>0.0076</b>	0.2487	<b>0.0076</b>	0.2457	0.0373	0.4121	0.0447	0.3721	0.0168	0.3624	0.0323	0.3947
	1:4	0.0269	0.3467	0.0243	0.3287	0.0227	0.3251	0.0729	0.4645	0.0713	0.4017	0.0484	0.4634	0.0639	0.4535
	1:9	0.0587	0.4085	0.0538	0.3794	0.0518	0.3896	0.1099	0.4883	0.0968	0.4097	0.1045	0.5212	0.1036	0.4819
10:30	1:1	<b>0.0090</b>	0.1908	0.0179	0.0765	<b>0.0077</b>	0.1820	<b>0.0139</b>	<b>0.7279</b>	0.0166	0.7239	<b>0.0117</b>	0.2185	<b>0.0094</b>	0.4672
	1:2.25	<b>0.0088</b>	0.2509	0.0041	0.2749	<b>0.0088</b>	0.2417	0.0700	0.9209	0.0748	0.8830	<b>0.0101</b>	<b>0.6492</b>	0.0212	0.6816
	1:4	<b>0.0090</b>	<b>0.2682</b>	0.0048	0.4055	<b>0.0073</b>	0.2601	0.1944	0.9522	0.1689	0.9116	0.0412	0.8277	0.0532	0.7432
	1:9	<b>0.0096</b>	<b>0.2826</b>	0.0168	0.5163	<b>0.0092</b>	0.2750	0.3510	0.9754	0.2644	0.9370	0.1348	0.9271	0.0972	0.7956
30:10	1:1	<b>0.0100</b>	0.1701	0.0185	0.3087	<b>0.0091</b>	0.1627	<b>0.0141</b>	0.3386	0.0164	0.3268	<b>0.0143</b>	0.3736	<b>0.0093</b>	<b>0.4910</b>
	1:2.25	0.0220	0.3990	0.0478	0.3807	0.0213	0.3886	0.0430	0.4329	0.0441	0.3910	0.0604	0.4604	0.0798	0.6268
	1:4	0.0532	0.5200	0.0640	0.4032	0.0508	0.5106	0.0739	0.4691	0.0696	0.4131	0.0945	0.4929	0.1464	0.6732
	1:9	0.1270	0.6063	0.0766	0.4024	0.1220	0.6002	0.0986	0.4845	0.0856	0.4123	0.1226	0.5180	0.2165	0.7136
30:30	1:1	0.0047	0.3356	0.0043	0.3308	0.0043	0.3260	<b>0.0093</b>	<b>0.8521</b>	<b>0.0103</b>	0.8308	0.0047	0.6516	<b>0.0087</b>	0.8481
	1:2.25	<b>0.0129</b>	<b>0.4988</b>	<b>0.0122</b>	0.4944	<b>0.0117</b>	0.4892	0.1223	0.9438	0.1154	0.9159	0.0435	0.8568	0.1184	0.9432
	1:4	0.0223	0.5557	0.0217	0.5477	0.0210	0.5512	0.2258	0.9629	0.1898	0.9382	0.1040	0.9134	0.2208	0.9627
	1:9	0.0424	0.5858	0.0417	0.5740	0.0403	0.5798	0.3577	0.9717	0.2828	0.9439	0.1996	0.9419	0.3550	0.9715
50:100	1:1	<b>0.0079</b>	0.6092	<b>0.0106</b>	0.6279	<b>0.0074</b>	0.6047	<b>0.0082</b>	<b>1.0000</b>	<b>0.0085</b>	0.9999	<b>0.0078</b>	0.9883	<b>0.0074</b>	0.9994
	1:2.25	0.0042	0.7058	<b>0.0063</b>	<b>0.8082</b>	0.0042	0.7016	0.3622	1.0000	0.3360	1.0000	0.0589	0.9998	0.2419	1.0000
	1:4	<b>0.0095</b>	0.7173	<b>0.0105</b>	<b>0.8439</b>	<b>0.0092</b>	0.7156	0.7481	1.0000	0.6806	1.0000	0.2411	1.0000	0.5775	1.0000
	1:9	<b>0.0084</b>	<b>0.7804</b>	0.0177	0.8860	<b>0.0086</b>	0.7571	0.9124	1.0000	0.8527	1.0000	0.5367	1.0000	0.7951	1.0000
100:50	1:1	<b>0.0077</b>	0.6244	<b>0.0099</b>	0.6187	<b>0.0074</b>	0.6191	<b>0.0111</b>	0.9887	<b>0.0118</b>	0.9861	<b>0.0086</b>	0.9382	<b>0.0103</b>	<b>0.9945</b>
	1:2.25	0.0215	0.7653	0.0222	0.6819	0.0209	0.7612	0.2116	0.9887	0.1986	0.9970	0.0805	0.9846	0.2911	0.9996
	1:4	0.0381	0.7925	0.0288	0.6861	0.0372	0.7911	0.4207	0.9993	0.3768	0.9985	0.1787	0.9930	0.5500	1.0000
	1:9	0.0778	0.8454	0.0424	0.7311	0.0755	0.8439	0.5922	0.9995	0.5275	0.9980	0.3278	0.9961	0.7315	1.0000
100:100	1:1	<b>0.0084</b>	0.7910	<b>0.0080</b>	0.7906	<b>0.0084</b>	0.7893	<b>0.0123</b>	<b>1.0000</b>	<b>0.0121</b>	<b>1.0000</b>	<b>0.0116</b>	0.9969	<b>0.0123</b>	<b>1.0000</b>
	1:2.25	<b>0.0104</b>	<b>0.8765</b>	<b>0.0103</b>	0.8752	<b>0.0107</b>	0.8730	0.4558	1.0000	0.4272	1.0000	0.1131	0.9999	0.4530	1.0000
	1:4	0.0202	0.8757	0.0201	0.8741	0.0199	0.8747	0.7719	1.0000	0.7214	1.0000	0.3060	1.0000	0.7691	1.0000
	1:9	0.0241	0.8805	0.0235	0.8781	0.0237	0.8779	0.9060	1.0000	0.8627	1.0000	0.5425	1.0000	0.9052	1.0000

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 2, it could be found that when the populations had log-normal distribution and equal variance at 0.01 significant level, WBR test and Exact Wilcoxon signed-rank test could control type I errors in all sample sizes. For unequal variance cases t test, t-BSE test and Welch t test could control type I errors in some sample sizes respectively.

It could be found that when the population has log-normal distribution and equal variance, the test statistics with the highest testing power that could control type I errors are WBR test (71.43%) and Exact Wilcoxon



signed-rank test(42.86%) respectively. For unequal variance, the test statistics with the highest testing power that could control were t test (28.57%) and Welch t test (9.52%) respectively.

**Table 3** The probability of a type I error and power test (0.05 significant level) of statistics on location testing between two populations with gamma distribution classified by sample size and variance

Sample Size	Var.ratio	t				Welch				t-BSE				WBR				B-M				Y-W				Wilcoxon			
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT																
10:10	1:1	<b>0.0440</b>	0.6170	<b>0.0881</b>	0.6109	<b>0.0484</b>	0.6269	<b>0.0531</b>	<b>0.7376</b>	<b>0.0564</b>	0.7056	<b>0.0393</b>	0.6090	<b>0.0448</b>	0.7134														
	1:2.25	<b>0.0546</b>	0.7129	<b>0.0469</b>	0.7003	<b>0.0589</b>	0.7199	0.0929	0.8046	0.0912	0.7344	<b>0.0643</b>	<b>0.7417</b>	0.0804	0.7859														
	1:4	<b>0.0718</b>	0.7485	<b>0.0686</b>	0.7328	<b>0.0750</b>	<b>0.7542</b>	0.1236	0.8256	0.1066	0.7285	0.0964	0.7765	0.1095	0.8104														
	1:9	0.0851	0.7739	0.0800	0.7541	0.0882	0.7789	0.1501	0.8322	0.1157	0.7112	0.1321	0.8062	0.1383	0.8249														
10:30	1:1	<b>0.0437</b>	0.7771	<b>0.0632</b>	0.8277	<b>0.0456</b>	0.7812	<b>0.0580</b>	<b>0.9809</b>	<b>0.0587</b>	0.9771	<b>0.0583</b>	0.8433	<b>0.0485</b>	0.9446														
	1:2.25	0.0238	0.8195	<b>0.0463</b>	0.9397	0.0248	0.8222	0.1169	0.9964	0.1132	0.9918	<b>0.0486</b>	<b>0.9788</b>	<b>0.0646</b>	0.9737														
	1:4	0.0187	0.8489	<b>0.0411</b>	<b>0.9677</b>	0.0197	0.8513	0.2187	0.9986	0.1933	0.9963	0.0771	0.9944	0.0970	0.9834														
	1:9	0.0164	0.8642	<b>0.0510</b>	<b>0.9803</b>	0.0163	0.8674	0.3386	0.9993	0.2645	0.9980	0.1517	0.9980	0.1367	0.9902														
30:10	1:1	<b>0.0458</b>	0.7807	<b>0.0622</b>	0.7195	<b>0.0463</b>	0.7885	<b>0.0512</b>	0.7630	<b>0.0640</b>	0.7443	<b>0.0564</b>	0.7097	<b>0.0477</b>	<b>0.8849</b>														
	1:2.25	0.1044	0.8705	0.0821	0.7378	0.1060	0.8719	0.0906	0.7980	0.0892	0.7617	0.1016	0.7593	0.1477	0.9066														
	1:4	0.1580	0.8977	0.0888	0.7484	0.1607	0.8992	0.1133	0.8123	0.1048	0.7626	0.1189	0.7773	0.2007	0.9129														
	1:9	0.2215	0.9132	0.0950	0.7554	0.2265	0.9145	0.1498	0.8176	0.1308	0.7347	0.1476	0.7973	0.2677	0.9160														
30:30	1:1	<b>0.0499</b>	0.9584	<b>0.0493</b>	0.9575	<b>0.0497</b>	0.9587	<b>0.0497</b>	<b>0.9946</b>	<b>0.0518</b>	0.9927	<b>0.0463</b>	0.9750	<b>0.0497</b>	<b>0.9946</b>														
	1:2.25	<b>0.0522</b>	0.9781	<b>0.0514</b>	0.9774	<b>0.0580</b>	<b>0.9788</b>	0.1586	0.9982	0.1492	0.9967	0.0786	0.9935	0.1586	0.9982														
	1:4	<b>0.0560</b>	<b>0.9888</b>	<b>0.0562</b>	0.9828	<b>0.0573</b>	0.9836	0.2526	0.9991	0.2177	0.9980	0.1266	0.9967	0.2527	0.9991														
	1:9	<b>0.0717</b>	0.9860	<b>0.0689</b>	0.9847	<b>0.0724</b>	<b>0.9884</b>	0.3617	0.9993	0.3006	0.9982	0.2001	0.9985	0.3626	0.9995														
50:100	1:1	<b>0.0484</b>	0.9999	<b>0.0468</b>	<b>1.0000</b>	<b>0.0496</b>	0.9999	<b>0.0494</b>	<b>1.0000</b>	<b>0.0489</b>	<b>1.0000</b>	<b>0.0467</b>	<b>1.0000</b>	<b>0.0479</b>	<b>1.0000</b>														
	1:2.25	<b>0.0293</b>	<b>1.0000</b>	<b>0.0483</b>	<b>1.0000</b>	<b>0.0292</b>	<b>1.0000</b>	0.3677	1.0000	0.3492	1.0000	0.1052	1.0000	0.2837	1.0000														
	1:4	0.0224	1.0000	<b>0.0494</b>	<b>1.0000</b>	0.0229	1.0000	0.6001	1.0000	0.5497	1.0000	0.2033	1.0000	0.4710	1.0000														
	1:9	0.0184	1.0000	<b>0.0546</b>	<b>1.0000</b>	0.0183	1.0000	0.7853	1.0000	0.7089	1.0000	0.3792	1.0000	0.6446	1.0000														
100:50	1:1	<b>0.0504</b>	0.9998	<b>0.0544</b>	0.9983	<b>0.0508</b>	0.9998	<b>0.0504</b>	<b>1.0000</b>	<b>0.0505</b>	<b>1.0000</b>	<b>0.0509</b>	0.9998	<b>0.0493</b>	<b>1.0000</b>														
	1:2.25	0.0913	1.0000	<b>0.0594</b>	<b>0.9994</b>	0.0925	1.0000	0.2378	1.0000	0.2259	1.0000	0.1053	1.0000	0.2970	1.0000														
	1:4	0.1127	0.9999	<b>0.0589</b>	<b>0.9995</b>	0.1130	0.9999	0.3967	1.0000	0.3680	1.0000	0.1718	1.0000	0.4877	1.0000														
	1:9	0.1424	1.0000	<b>0.0631</b>	<b>0.9996</b>	0.1435	1.0000	0.4986	1.0000	0.4502	1.0000	0.2632	1.0000	0.6082	1.0000														
100:100	1:1	<b>0.0484</b>	<b>1.0000</b>	<b>0.0481</b>	<b>1.0000</b>	<b>0.0490</b>	<b>1.0000</b>	<b>0.0494</b>	<b>1.0000</b>	<b>0.0497</b>	<b>1.0000</b>	<b>0.0478</b>	<b>1.0000</b>	<b>0.0494</b>	<b>1.0000</b>														
	1:2.25	<b>0.0510</b>	<b>1.0000</b>	<b>0.0507</b>	<b>1.0000</b>	<b>0.0511</b>	<b>1.0000</b>	0.3964	1.0000	0.3756	1.0000	0.1207	1.0000	0.3948	1.0000														
	1:4	<b>0.0570</b>	<b>1.0000</b>	<b>0.0568</b>	<b>1.0000</b>	<b>0.0582</b>	<b>1.0000</b>	0.6616	1.0000	0.6210	1.0000	0.2539	1.0000	0.6614	1.0000														
	1:9	<b>0.0574</b>	<b>1.0000</b>	<b>0.0569</b>	<b>1.0000</b>	<b>0.0578</b>	<b>1.0000</b>	0.8008	1.0000	0.7401	1.0000	0.4189	1.0000	0.8008	1.0000														

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 3, it could be found that when the populations had gamma distribution and equal variance at 0.05 significant level, all test statistics could control type I errors in all sample sizes. For unequal variance cases Welch t test, t test and t-BSE test could control type I errors in some sample sizes respectively.

It could be found that when the population has gamma distribution and equal variance, the test statistics with the highest testing power that could control type I errors are WBR test (85.71%) and Exact Wilcoxon signed-rank test (71.43%) respectively. For unequal variance, the test statistics with the highest testing power that could control type I errors were Welch t test (52.38%), t-BSE test (33.33%) and t test (23.81%) respectively.



**Table 4** The probability of a type I error and power test (0.01 significant level) of statistics on location testing between two populations with gamma distribution classified by sample size and variance

Sample Size	Var.ratio	t		Welch		t-BSE		WBR		B-M		Y-W		Wilcoxon	
		TE	PT												
10:10	1:1	<b>0.0068</b>	0.3737	0.0042	0.3553	0.0048	0.3417	<b>0.0118</b>	<b>0.4951</b>	0.0167	0.4900	<b>0.0068</b>	0.3132	<b>0.0080</b>	0.4635
	1:2.25	<b>0.0127</b>	<b>0.5221</b>	<b>0.0103</b>	0.5043	<b>0.0100</b>	0.4926	0.0248	0.5793	0.0314	0.5334	<b>0.0138</b>	0.5071	0.0200	0.5599
	1:4	0.0221	0.5744	0.0202	0.5465	0.0190	0.5473	0.0409	0.6009	0.0428	0.5320	0.0281	0.5716	0.0360	0.5879
	1:9	0.0351	0.6186	0.0324	0.5773	0.0311	0.5961	0.0536	0.6213	0.0494	0.5298	0.0523	0.5996	0.0497	0.6151
10:30	1:1	<b>0.0068</b>	0.5563	0.0218	0.5019	<b>0.0078</b>	0.5404	<b>0.0140</b>	<b>0.9299</b>	0.0163	0.9245	0.0167	0.4981	0.0087	0.7703
	1:2.25	<b>0.0069</b>	0.6136	<b>0.0093</b>	0.8044	<b>0.0064</b>	0.6012	0.0370	0.9785	0.0402	0.9658	<b>0.0068</b>	0.8715	<b>0.0124</b>	<b>0.8733</b>
	1:4	<b>0.0065</b>	0.6475	<b>0.0076</b>	<b>0.8965</b>	<b>0.0068</b>	0.6332	0.0825	0.9899	0.0775	0.9775	0.0186	0.9578	0.0178	0.8986
	1:9	0.0041	0.6737	<b>0.0101</b>	<b>0.9812</b>	0.0037	0.6591	0.1612	0.9958	0.1137	0.9838	0.0546	0.9872	0.0270	0.9171
30:10	1:1	<b>0.0085</b>	0.5741	0.0181	0.5435	<b>0.0077</b>	0.5597	<b>0.0147</b>	0.5415	0.0178	0.5204	<b>0.0147</b>	0.5185	<b>0.0090</b>	<b>0.8961</b>
	1:2.25	0.0276	0.7602	0.0329	0.5630	0.0250	0.7542	0.0282	0.5906	0.0290	0.5430	0.0423	0.5570	0.0489	0.7679
	1:4	0.0605	0.8166	0.0394	0.5766	0.0575	0.8102	0.0404	0.6017	0.0375	0.5357	0.0545	0.5688	0.0851	0.7912
	1:9	0.1094	0.8535	0.0445	0.5743	0.1044	0.8492	0.0591	0.6067	0.0520	0.5237	0.0710	0.5824	0.1314	0.8090
30:30	1:1	<b>0.0088</b>	0.8731	<b>0.0073</b>	0.8716	<b>0.0078</b>	0.8669	<b>0.0106</b>	<b>0.9751</b>	<b>0.0114</b>	0.9686	<b>0.0068</b>	0.9083	<b>0.0097</b>	0.9736
	1:2.25	<b>0.0126</b>	<b>0.9309</b>	<b>0.0118</b>	0.9271	<b>0.0121</b>	0.9278	0.0544	0.9893	0.0526	0.9831	0.0222	0.9699	0.0514	0.9889
	1:4	0.0174	0.9462	0.0168	0.9411	0.0167	0.9437	0.1053	0.9935	0.0877	0.9869	0.0472	0.9837	0.1021	0.9933
	1:9	0.0236	0.9572	0.0223	0.9520	0.0230	0.9552	0.1772	0.9943	0.1284	0.9869	0.0887	0.9902	0.1754	0.9942
50:100	1:1	<b>0.0089</b>	0.9979	<b>0.0103</b>	0.9994	<b>0.0089</b>	0.9980	<b>0.0087</b>	<b>1.0000</b>	<b>0.0089</b>	<b>1.0000</b>	<b>0.0084</b>	<b>1.0000</b>	<b>0.0084</b>	<b>1.0000</b>
	1:2.25	0.0048	0.9999	<b>0.0091</b>	<b>1.0000</b>	<b>0.0060</b>	0.9999	0.1728	1.0000	0.1599	1.0000	0.0262	1.0000	0.1070	1.0000
	1:4	0.0039	1.0000	<b>0.0118</b>	<b>1.0000</b>	0.0041	0.9999	0.3693	1.0000	0.3149	1.0000	0.0786	1.0000	0.2265	1.0000
	1:9	0.0028	0.9999	<b>0.0182</b>	<b>1.0000</b>	0.0026	0.9999	0.5880	1.0000	0.4730	1.0000	0.1940	1.0000	0.3760	1.0000
100:50	1:1	<b>0.0088</b>	0.9966	<b>0.0111</b>	0.9919	<b>0.0089</b>	0.9965	<b>0.0100</b>	<b>0.9998</b>	<b>0.0108</b>	<b>0.9998</b>	<b>0.0099</b>	0.9979	<b>0.0094</b>	<b>0.9998</b>
	1:2.25	0.0250	0.9992	0.0177	0.9956	0.0259	0.9990	0.0931	1.0000	0.0869	0.9999	0.0363	0.9994	0.1415	1.0000
	1:4	0.0390	0.9997	0.0193	0.9968	0.0383	0.9996	0.1918	1.0000	0.1641	0.9999	0.0742	0.9999	0.2829	1.0000
	1:9	0.0602	0.9997	0.0189	0.9974	0.0588	0.9997	0.2807	0.9999	0.2333	0.9997	0.1267	0.9997	0.4125	1.0000
100:100	1:1	<b>0.0083</b>	<b>1.0000</b>	<b>0.0083</b>	<b>1.0000</b>	<b>0.0084</b>	<b>1.0000</b>	<b>0.0099</b>	<b>1.0000</b>	<b>0.0102</b>	<b>1.0000</b>	<b>0.0081</b>	<b>1.0000</b>	<b>0.0095</b>	<b>1.0000</b>
	1:2.25	<b>0.0105</b>	<b>1.0000</b>	<b>0.0104</b>	<b>1.0000</b>	<b>0.0105</b>	1.0000	0.1943	1.0000	0.1787	1.0000	0.0393	1.0000	0.1916	1.0000
	1:4	<b>0.0140</b>	<b>1.0000</b>	<b>0.0138</b>	<b>1.0000</b>	<b>0.0137</b>	1.0000	0.4350	1.0000	0.3793	1.0000	0.1111	1.0000	0.4316	1.0000
	1:9	0.0157	1.0000	0.0156	1.0000	0.0151	1.0000	0.6054	1.0000	0.5130	1.0000	0.2335	1.0000	0.6033	1.0000

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 4, it could be found that when the populations had gamma distribution and equal variance at 0.01 significant level, The t test, Welch t test and Exact Wilcoxon signed-rank test could control type I errors in all sample sizes. For unequal variance cases Welch t test, t-BSE and t test could control type I errors in some sample sizes respectively

It could be found that when the population had gamma distribution and equal variance, the test statistics with the highest testing power that could control type I errors are WBR test(85.71%) and Exact Wilcoxon signed-rank test(57.14%) respectively. For unequal variance, the test statistics with the highest testing power that can control type I errors were Welch t test (33.33%) and t-test (19.05%) respectively.



**Table 5** The probability of a type I error and power test (0.05 significant level) of statistics on location testing between two populations with exponential distribution classified by sample size and variance

Sample Size	Var.ratio	t				Welch				t-BSE				WBR				B-M				Y-W				Wilcoxon	
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT														
10:10	1:1	<b>0.0401</b>	0.6122	<b>0.0349</b>	0.6037	<b>0.0446</b>	0.6224	<b>0.0492</b>	<b>0.7330</b>	<b>0.0525</b>	0.6972	<b>0.0362</b>	0.6034	<b>0.0403</b>	0.7054												
	1:2.25	<b>0.0550</b>	0.7163	<b>0.0508</b>	0.7041	<b>0.0582</b>	0.7213	0.0927	0.8042	0.0887	0.7353	<b>0.0642</b>	<b>0.7376</b>	0.0781	0.7836												
	1:4	<b>0.0699</b>	0.7497	<b>0.0673</b>	0.7338	<b>0.0740</b>	<b>0.7558</b>	0.1215	0.8218	0.1059	0.7323	0.0943	0.7779	0.1076	0.8076												
10:30	1:9	0.0895	0.7754	0.0844	0.7545	0.0948	0.7807	0.1550	0.8292	0.1188	0.6989	0.1331	0.8011	0.1438	0.8215												
	1:1	<b>0.0471</b>	0.7769	<b>0.0689</b>	0.8228	<b>0.0500</b>	0.7793	<b>0.0564</b>	<b>0.9834</b>	<b>0.0588</b>	0.9786	<b>0.0613</b>	0.8476	<b>0.0523</b>	0.9464												
	1:2.25	<b>0.0277</b>	0.8171	<b>0.0466</b>	0.9362	<b>0.0293</b>	0.8195	0.1263	0.9961	0.1251	0.9926	<b>0.0488</b>	<b>0.9783</b>	<b>0.0678</b>	0.9732												
	1:4	0.0187	0.8535	<b>0.0443</b>	<b>0.9664</b>	0.0195	0.8556	0.2331	0.9985	0.2056	0.9948	0.0801	0.9923	0.1050	0.9829												
30:10	1:9	0.0146	0.8712	<b>0.0510</b>	<b>0.9814</b>	0.0146	0.8750	0.3301	0.9993	0.2594	0.9972	0.1475	0.9980	0.1304	0.9909												
	1:1	<b>0.0446</b>	0.7867	<b>0.0648</b>	0.7263	<b>0.0462</b>	0.7895	<b>0.0570</b>	0.7734	<b>0.0606</b>	0.7553	<b>0.0629</b>	0.7137	<b>0.0531</b>	<b>0.8624</b>												
	1:2.25	0.1023	0.8723	0.0837	0.7429	0.1051	0.8738	0.0881	0.8042	0.0855	0.7675	0.0982	0.7639	0.1414	0.9008												
	1:4	0.1582	0.9012	0.0895	0.7582	0.1610	0.9018	0.1169	0.8141	0.1076	0.7646	0.1226	0.7841	0.2071	0.9155												
30:30	1:9	0.2286	0.9172	0.0932	0.7537	0.2301	0.9191	0.1471	0.8226	0.1235	0.7429	0.1474	0.8012	0.2638	0.9203												
	1:1	<b>0.0487</b>	0.9504	<b>0.0478</b>	0.9497	<b>0.0496</b>	0.9520	<b>0.0524</b>	<b>0.9937</b>	<b>0.0531</b>	0.9923	<b>0.0476</b>	0.9752	<b>0.0524</b>	<b>0.9937</b>												
	1:2.25	<b>0.0534</b>	<b>0.9771</b>	<b>0.0527</b>	0.9761	<b>0.0558</b>	0.9770	0.1669	0.9986	0.1583	0.9975	0.0823	0.9937	0.1669	0.9986												
	1:4	<b>0.0612</b>	<b>0.9846</b>	<b>0.0601</b>	0.9830	<b>0.0625</b>	0.9838	0.2625	0.9993	0.2296	0.9986	0.1235	0.9976	0.2625	0.9993												
50:100	1:9	<b>0.0693</b>	<b>0.9847</b>	<b>0.0665</b>	0.9833	<b>0.0711</b>	<b>0.9847</b>	0.3629	0.9991	0.3010	0.9975	0.2021	0.9974	0.3640	0.9991												
	1:1	<b>0.0465</b>	0.9996	<b>0.0493</b>	0.9998	<b>0.0478</b>	0.9996	<b>0.0525</b>	<b>1.0000</b>	<b>0.0529</b>	<b>1.0000</b>	<b>0.0494</b>	<b>1.0000</b>	<b>0.0519</b>	<b>1.0000</b>												
	1:2.25	<b>0.0266</b>	0.9999	<b>0.0475</b>	<b>1.0000</b>	<b>0.0272</b>	0.9999	0.3508	1.0000	0.3300	1.0000	0.0963	1.0000	0.2707	1.0000												
	1:4	0.0200	1.0000	<b>0.0517</b>	<b>1.0000</b>	0.0202	1.0000	0.6078	1.0000	0.5548	1.0000	0.1994	1.0000	0.4739	1.0000												
100:50	1:9	0.0157	1.0000	<b>0.0532</b>	<b>1.0000</b>	0.0158	1.0000	0.7895	1.0000	0.7132	1.0000	0.3822	1.0000	0.6528	1.0000												
	1:1	<b>0.0404</b>	0.9994	<b>0.0471</b>	0.9988	<b>0.0403</b>	0.9995	<b>0.0442</b>	<b>1.0000</b>	<b>0.0452</b>	<b>1.0000</b>	<b>0.0463</b>	<b>1.0000</b>	<b>0.0442</b>	<b>1.0000</b>												
	1:2.25	0.0876	1.0000	<b>0.0594</b>	<b>0.9997</b>	0.0888	1.0000	0.2545	1.0000	0.2437	1.0000	0.1153	1.0000	0.3170	1.0000												
	1:4	0.1208	0.9999	<b>0.0676</b>	<b>0.9994</b>	0.1222	1.0000	0.4120	1.0000	0.3831	1.0000	0.1928	1.0000	0.5051	1.0000												
100:100	1:9	0.1472	1.0000	<b>0.0651</b>	<b>0.9995</b>	0.1477	1.0000	0.4924	1.0000	0.4451	1.0000	0.2534	1.0000	0.6060	1.0000												
	1:1	<b>0.0496</b>	<b>1.0000</b>	<b>0.0494</b>	<b>1.0000</b>	<b>0.0502</b>	<b>1.0000</b>	<b>0.0474</b>	<b>1.0000</b>	<b>0.0479</b>	<b>1.0000</b>	<b>0.0470</b>	<b>1.0000</b>	<b>0.0473</b>	<b>1.0000</b>												
	1:2.25	<b>0.0516</b>	<b>1.0000</b>	<b>0.0514</b>	<b>1.0000</b>	<b>0.0538</b>	<b>1.0000</b>	<b>0.0540</b>	<b>1.0000</b>	<b>0.0540</b>	<b>1.0000</b>	<b>0.0540</b>	<b>1.0000</b>	<b>0.0540</b>	<b>1.0000</b>												
	1:4	<b>0.0533</b>	<b>1.0000</b>	<b>0.0528</b>	<b>1.0000</b>	<b>0.0540</b>	<b>1.0000</b>																				
100:100	1:9	<b>0.0578</b>	<b>1.0000</b>	<b>0.0571</b>	<b>1.0000</b>	<b>0.0583</b>	<b>1.0000</b>																				

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 5, it could be found that when the populations have exponential distribution and equal variance at 0.05 significant level, all test statistics can control type I errors in all sample sizes. For unequal variance cases Welch t test can control type I errors in almost all sample sizes.

It could be found that when the population has exponential distribution and equal variance, the test statistics with the highest testing power that can control type I errors are WBR test (85.71%), and Exact Wilcoxon signed-rank test(71.43%)respectively. For unequal variance, the test statistics with the highest testing power that could control were Welch t test (52.38%).

**Table 6** The probability of a type I error and power test (0.01 significant level) of statistics on location testing between two populations which have exponential distribution classified by sample size and variance

Sample Size	Var.ratio	t				Welch				t-BSE				WBR				B-M				Y-W				Wilcoxon	
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT		
10:10	1:1	<b>0.0057</b>	0.3708	0.0035	0.3524	0.0037	0.3365	<b>0.0096</b>	0.4905	<b>0.0142</b>	<b>0.4906</b>	<b>0.0060</b>	0.3102	<b>0.0071</b>	0.4606												
	1:2.25	<b>0.0122</b>	<b>0.5215</b>	<b>0.0094</b>	0.5028	<b>0.0100</b>	0.4931	0.0248	0.5769	0.0318	0.5335	<b>0.0139</b>	0.5077	0.0197	0.5561												
	1:4	0.0218	0.5793	0.0193	0.5516	0.0181	0.5557	0.0409	0.6039	0.0431	0.5281	0.0293	0.5688	0.0362	0.5941												
10:30	1:9	0.0333	0.6243	0.0302	0.5791	0.0289	0.6003	0.0530	0.6164	0.0471	0.5221	0.0530	0.6052	0.0483	0.6087												
	1:1	<b>0.0094</b>	0.5627	<b>0.0218</b>	0.4975	<b>0.0088</b>	0.5463	<b>0.0150</b>	<b>0.9300</b>	0.0180	0.9252	0.0164	0.9479	<b>0.0102</b>	0.7682												
	1:2.25	<b>0.0063</b>	0.6129	<b>0.0078</b>	0.7979	<b>0.0059</b>	0.5987	0.0339	0.9772	0.0407	0.9645	<b>0.0107</b>	0.8690	<b>0.0110</b>	<b>0.8636</b>												
30:10	1:4	0.0040	0.6592	<b>0.0071</b>	<b>0.8943</b>	0.0039	0.6450	0.0861	0.9891	0.0799	0.9771	0.0174	0.9587	0.0175	0.9004												
	1:9	0.0023	0.6878	<b>0.0105</b>	0.9372	0.0019	0.6749	0.0952	0.9722	0.0172	0.9854	0.0499	0.9877	0.0278	0.9229												
	1:1	<b>0.008</b>																									



From table 6, it could be found that when the populations had exponential distribution and equal variance at 0.01 significant level, t test, WBR test and Exact Wilcoxon signed-rank test could control type I errors in all sample sizes. For unequal variance cases Welch t test, t test and t-BSE test can control type I errors in some sample sizes respectively.

It could be found that when a population had exponential distribution and equal variance, the test statistics with the highest testing power that could control type I errors are Exact Wilcoxon signed-rank test (71.43%), WBR test (42.86%) and B-M test (42.86%) respectively. For unequal variance, the test statistics with the highest testing power that could control were Welch t test (23.81%) and t test (19.05%) respectively.

**Table 7** The probability of a type I error and power test (0.05 significant level) of statistics on location testing between two populations with poisson distribution classified by sample size and variance

Sample Size	Var.ratio	t		Welch		t-BSE		WBR		B-M		Y-W		Wilcoxon	
		TE	PT	TE	PT	TE	PT								
10:10	1:1	<b>0.0493</b>	0.1558	<b>0.0478</b>	0.1522	<b>0.0533</b>	0.1637	<b>0.0511</b>	0.1588	<b>0.0569</b>	<b>0.1728</b>	<b>0.0469</b>	0.1286	<b>0.0447</b>	0.1433
	1:2.25	<b>0.0516</b>	0.2108	<b>0.0489</b>	0.2052	<b>0.0553</b>	0.2215	<b>0.0558</b>	<b>0.2258</b>	<b>0.0571</b>	0.2234	<b>0.0523</b>	0.1772	<b>0.0493</b>	0.2034
	1:4	<b>0.0567</b>	0.2586	<b>0.0522</b>	0.2488	<b>0.0591</b>	0.2659	<b>0.0635</b>	<b>0.2744</b>	<b>0.0590</b>	0.2545	<b>0.0549</b>	0.2209	<b>0.0554</b>	0.2553
	1:9	<b>0.0602</b>	0.2865	<b>0.0526</b>	0.2652	<b>0.0641</b>	0.2943	<b>0.0716</b>	<b>0.3013</b>	<b>0.0562</b>	0.2535	<b>0.0608</b>	0.2362	<b>0.0648</b>	0.2865
10:30	1:1	<b>0.0481</b>	0.2130	<b>0.0525</b>	0.1816	<b>0.0504</b>	0.2169	<b>0.0526</b>	0.2127	<b>0.0552</b>	<b>0.2221</b>	<b>0.0578</b>	0.1682	<b>0.0447</b>	0.1938
	1:2.25	0.0181	0.2252	<b>0.0504</b>	0.3327	0.0186	0.2300	<b>0.0609</b>	0.3895	<b>0.0599</b>	<b>0.3901</b>	<b>0.0581</b>	0.3075	<b>0.0324</b>	0.2844
	1:4	0.0107	0.2088	<b>0.0485</b>	0.4197	0.0113	0.2130	<b>0.0683</b>	<b>0.5134</b>	<b>0.0603</b>	0.4916	<b>0.0528</b>	0.3973	0.0224	0.3220
	1:9	0.0047	0.2034	<b>0.0528</b>	0.5461	0.0050	0.2082	0.0911	0.6484	<b>0.0636</b>	<b>0.5760</b>	<b>0.0578</b>	0.5333	0.0207	0.3775
30:10	1:1	<b>0.0482</b>	0.2283	<b>0.0520</b>	<b>0.2364</b>	<b>0.0496</b>	0.2324	<b>0.0534</b>	0.2129	<b>0.0562</b>	0.2229	<b>0.0585</b>	0.2060	<b>0.0464</b>	0.2143
	1:2.25	0.1050	0.4025	<b>0.0525</b>	0.2671	0.1072	0.4067	<b>0.0567</b>	<b>0.2698</b>	<b>0.0561</b>	0.2657	<b>0.0592</b>	0.2336	0.0779	0.3464
	1:4	0.1522	0.4728	<b>0.0548</b>	0.2672	0.1538	0.4763	<b>0.0597</b>	<b>0.2824</b>	<b>0.0570</b>	0.2679	<b>0.0627</b>	0.2285	0.0994	0.3848
	1:9	0.2154	0.5533	<b>0.0532</b>	0.2707	0.2206	0.5573	<b>0.0666</b>	<b>0.2964</b>	<b>0.0597</b>	0.2684	<b>0.0605</b>	0.2374	0.1357	0.4360
30:30	1:1	<b>0.0493</b>	0.4037	<b>0.0493</b>	0.4032	<b>0.0511</b>	<b>0.4073</b>	<b>0.0493</b>	0.3934	<b>0.0510</b>	0.4022	<b>0.0504</b>	0.3403	<b>0.0476</b>	0.3875
	1:2.25	<b>0.0503</b>	0.5056	<b>0.0497</b>	0.5023	<b>0.0517</b>	0.5104	<b>0.0562</b>	<b>0.5505</b>	<b>0.0531</b>	0.5383	<b>0.0525</b>	0.4656	<b>0.0543</b>	0.5461
	1:4	<b>0.0499</b>	0.5930	<b>0.0485</b>	0.5865	<b>0.0501</b>	0.5965	<b>0.0633</b>	<b>0.6464</b>	<b>0.0555</b>	0.6150	<b>0.0521</b>	0.5505	<b>0.0620</b>	0.6425
	1:9	<b>0.0512</b>	0.6373	<b>0.0491</b>	0.6271	<b>0.0534</b>	0.6387	0.0807	0.6964	<b>0.0614</b>	0.6393	<b>0.0560</b>	0.6097	0.0797	0.6946
50:100	1:1	<b>0.0539</b>	0.7265	<b>0.0531</b>	0.7269	<b>0.0548</b>	0.7275	<b>0.0618</b>	0.7294	<b>0.0593</b>	<b>0.7328</b>	<b>0.0574</b>	0.6520	<b>0.0606</b>	0.7213
	1:2.25	<b>0.0268</b>	0.8294	<b>0.0490</b>	0.8951	<b>0.0272</b>	0.8301	<b>0.0601</b>	<b>0.9300</b>	<b>0.0558</b>	0.9250	<b>0.0545</b>	0.8639	<b>0.0439</b>	0.9009
	1:4	0.0180	0.8706	<b>0.0537</b>	0.9466	0.0177	0.8719	0.0834	0.9720	<b>0.0703</b>	<b>0.9646</b>	<b>0.0635</b>	0.9362	<b>0.0481</b>	0.9490
	1:9	0.0108	0.9215	<b>0.0471</b>	<b>0.9794</b>	0.0110	0.9213	0.1224	0.9907	0.0869	0.9836	0.0760	0.9794	<b>0.0617</b>	0.9766
100:50	1:1	0.1052	0.7200	<b>0.0537</b>	<b>0.7097</b>	0.1072	0.7197	<b>0.0568</b>	0.6991	<b>0.0565</b>	0.7018	<b>0.0630</b>	0.6318	0.0795	0.7091
	1:2.25	0.0860	0.8558	<b>0.0519</b>	0.7880	0.0872	0.8571	<b>0.0625</b>	<b>0.8236</b>	<b>0.0593</b>	0.8157	<b>0.0555</b>	0.7441	0.0792	0.8625
	1:4	0.1158	0.8963	<b>0.0531</b>	0.8059	0.1153	0.8963	<b>0.0677</b>	<b>0.8503</b>	<b>0.0605</b>	0.8327	<b>0.0582</b>	0.7831	0.1042	0.8980
	1:9	0.1384	0.9338	<b>0.0521</b>	0.8430	0.1379	0.9338	0.0870	0.8748	<b>0.0719</b>	0.8471	<b>0.0666</b>	0.8377	0.1417	0.9230
100:100	1:1	<b>0.0287</b>	<b>0.8730</b>	<b>0.0530</b>	<b>0.8730</b>	<b>0.0278</b>	0.8727	<b>0.0619</b>	0.8647	<b>0.0577</b>	0.8662	<b>0.0564</b>	0.7993	<b>0.0440</b>	0.8644
	1:2.25	<b>0.0504</b>	0.9498	<b>0.0503</b>	0.9495	<b>0.0505</b>	0.9505	<b>0.0667</b>	<b>0.9703</b>	<b>0.0623</b>	0.9665	<b>0.0593</b>	0.9336	<b>0.0665</b>	0.9702
	1:4	<b>0.0527</b>	0.9795	<b>0.0523</b>	0.9792	<b>0.0534</b>	<b>0.9796</b>	0.0992	0.9898	0.0879	0.9856	0.0739	0.9728	0.0987	0.9898
	1:9	<b>0.0502</b>	<b>0.9829</b>	<b>0.0494</b>	0.9826	<b>0.0509</b>	0.9828	0.1112	0.9923	0.0815	0.9878	0.0754	0.9851	0.1111	0.9923

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 7, it could be found that when the populations have poisson distribution and equal variance at 0.05 significant level, Welch t test, WBR test, B-M test and Y-W test can control type I errors in all sample sizes. For unequal variance cases Welch t test, could control type I errors in all sample sizes followed by Y-W test, B-M test, WBR test respectively.

It could be found that when a population had poisson distribution and equal variance, the test statistics with the highest testing power that could control type I errors were Welch t test (42.86%) and WBR test (42.86%). For unequal variance, the test statistics with the highest testing power that could control almost all sample sizes was WBR test (61.90%).



**Table 8** The probability of a type I error and power test (0.01 significant level) of statistics on location testing between two populations with poisson distribution classified by sample size and variance

Sample Size	Var.ratio	t				Welch				t-BSE				WBR				B-M				Y-W				Wilcoxon	
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT		
10:10	1:1	<b>0.0093</b>	0.0500	<b>0.0089</b>	0.0468	0.0076	0.0412	<b>0.0103</b>	<b>0.0535</b>	0.0173	0.0733	<b>0.0102</b>	0.0325	<b>0.0073</b>	0.0349												
	1:2.25	<b>0.0107</b>	0.0750	<b>0.0101</b>	0.0708	<b>0.0092</b>	0.0656	<b>0.0146</b>	<b>0.0837</b>	0.0185	0.0986	<b>0.0101</b>	0.0569	<b>0.0086</b>	0.0598												
	1:4	<b>0.0134</b>	0.1047	<b>0.0115</b>	0.0950	<b>0.0109</b>	0.0887	<b>0.0150</b>	<b>0.1152</b>	0.0195	0.1228	<b>0.0135</b>	0.0802	<b>0.0092</b>	0.0830												
	1:9	0.0164	0.1322	<b>0.0124</b>	0.1114	<b>0.0136</b>	0.1196	0.0207	0.1421	0.0190	0.1304	0.0184	0.1010	<b>0.0135</b>	0.1127												
10:30	1:1	<b>0.0092</b>	0.0768	<b>0.0107</b>	0.0568	<b>0.0081</b>	0.0719	<b>0.0125</b>	<b>0.0856</b>	0.0163	0.0990	0.0155	0.0570	<b>0.0079</b>	0.0592												
	1:2.25	<b>0.0014</b>	0.0732	<b>0.0092</b>	0.1380	0.0011	0.0680	<b>0.0139</b>	0.1906	<b>0.0142</b>	<b>0.2023</b>	<b>0.0135</b>	0.1214	0.0028	0.0931												
	1:4	0.0007	0.0602	<b>0.0097</b>	<b>0.1973</b>	0.0006	0.0547	0.0154	0.2774	0.0155	0.2716	<b>0.0126</b>	0.1732	0.0028	0.1026												
	1:9	0.0001	0.0515	<b>0.0106</b>	<b>0.3013</b>	0.0001	0.0484	0.0242	0.4127	0.0158	0.3410	<b>0.0139</b>	0.2779	0.0024	0.1272												
30:10	1:1	<b>0.0086</b>	0.0752	<b>0.0106</b>	<b>0.0116</b>	<b>0.0077</b>	0.0713	<b>0.0122</b>	0.0809	0.0156	0.0962	<b>0.0146</b>	0.0787	<b>0.0072</b>	0.0671												
	1:2.25	0.0312	0.2209	<b>0.0117</b>	0.1123	0.0292	0.2118	<b>0.0139</b>	<b>0.1227</b>	0.0163	0.1259	0.0175	0.1032	0.0178	0.1610												
	1:4	0.0599	0.3036	<b>0.0141</b>	<b>0.1087</b>	0.0573	0.2938	0.0194	0.1288	0.0204	0.1261	0.0219	0.1048	0.0301	0.1996												
	1:9	0.1012	0.3971	<b>0.0134</b>	<b>0.1153</b>	0.0974	0.3874	0.0224	0.1451	0.0200	0.1313	0.0225	0.1072	0.0462	0.2500												
30:30	1:1	<b>0.0097</b>	0.1809	<b>0.0094</b>	0.1801	<b>0.0094</b>	0.1748	<b>0.0110</b>	0.1821	<b>0.0128</b>	<b>0.1993</b>	<b>0.0117</b>	0.1429	<b>0.0097</b>	0.1675												
	1:2.25	<b>0.0114</b>	0.2673	<b>0.0113</b>	0.2630	<b>0.0103</b>	0.2611	<b>0.0137</b>	<b>0.3103</b>	<b>0.0138</b>	<b>0.3076</b>	<b>0.0111</b>	0.2332	<b>0.0121</b>	0.2929												
	1:4	<b>0.0118</b>	0.3518	<b>0.0110</b>	<b>0.0111</b>	0.3450	0.0164	0.4073	<b>0.0147</b>	0.3795	<b>0.0144</b>	0.3104	<b>0.0147</b>	<b>0.3922</b>													
	1:9	<b>0.0135</b>	<b>0.4031</b>	<b>0.0118</b>	0.3833	<b>0.0125</b>	0.3940	0.0223	0.4707	0.0154	0.4003	<b>0.0150</b>	0.3725	0.0206	0.4589												
50:100	1:1	<b>0.0111</b>	0.4913	<b>0.0108</b>	0.4836	<b>0.0105</b>	0.4865	0.0152	0.4986	<b>0.0150</b>	<b>0.5113</b>	<b>0.0136</b>	0.3927	<b>0.0140</b>	0.4782												
	1:2.25	0.0034	0.6090	<b>0.0093</b>	0.7375	0.0035	0.6046	<b>0.0138</b>	<b>0.8084</b>	<b>0.0129</b>	0.7998	<b>0.0130</b>	0.6755	<b>0.0072</b>	0.7336												
	1:4	0.0014	0.6573	<b>0.0097</b>	<b>0.8345</b>	0.0014	0.6521	0.0227	0.9066	0.0183	0.8860	0.0154	0.8013	<b>0.0074</b>	0.8213												
	1:9	0.0002	0.7368	<b>0.0100</b>	<b>0.9211</b>	0.0002	0.7285	0.0390	0.9646	0.0230	0.9383	0.0190	0.9159	<b>0.0118</b>	0.9049												
100:50	1:1	0.0339	0.4928	<b>0.0115</b>	<b>0.4877</b>	0.0316	0.4886	<b>0.0147</b>	0.4657	0.0166	0.4761	<b>0.0191</b>	0.3962	0.0184	0.4741												
	1:2.25	0.0244	0.6997	<b>0.0122</b>	0.5757	0.0242	0.6976	0.0153	0.6217	<b>0.0141</b>	<b>0.6089</b>	<b>0.0131</b>	0.4996	0.0227	0.6937												
	1:4	0.0392	0.7763	<b>0.0107</b>	<b>0.6021</b>	0.0377	0.7733	0.0183	0.6647	0.0156	0.6359	<b>0.0132</b>	0.5485	0.0323	0.7607												
	1:9	0.0539	0.8517	<b>0.0111</b>	<b>0.6442</b>	0.0524	0.8482	0.0237	0.7139	0.0178	0.6630	<b>0.0148</b>	0.6126	0.0528	0.8189												
100:100	1:1	0.0027	0.6930	<b>0.0111</b>	0.6930	0.0027	0.6888	<b>0.0131</b>	0.6852	<b>0.0124</b>	<b>0.6938</b>	<b>0.0131</b>	0.5751	<b>0.0075</b>	0.6801												
	1:2.25	<b>0.0094</b>	0.8517	<b>0.0094</b>	0.8507	<b>0.0091</b>	0.8492	0.0153	0.8954	<b>0.0136</b>	<b>0.8873</b>	<b>0.0134</b>	0.8047	0.0151	0.8931												
	1:4	<b>0.0111</b>	<b>0.9200</b>	<b>0.0109</b>	0.9185	<b>0.0105</b>	0.9180	0.0287	0.9559	0.0230	0.9441	0.0205	0.8984	0.0280	0.9547												
	1:9	<b>0.0109</b>	<b>0.9388</b>	<b>0.0105</b>	0.9371	<b>0.0106</b>	0.9343	0.0243	0.9677	0.0202	0.9515	0.0194	0.9358	0.0337	0.9672												

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 8, it could be found that when the populations had poisson distribution and equal variance at 0.01 significant level, Welch t test could control type I errors in all sample sizes followed by WBR test, Y-W test and Exact Wilcoxon signed-rank test. For unequal variance cases Welch t test could control type I errors in all sample sizes.

It could be found that when the population has poisson distribution and equal variance, the test statistics with the highest testing power that could control type I errors was B-M test (42.86%). For unequal variance, the test statistics with the highest testing power that could control type I errors was Welch t test (38.10%).

**Table 9** The probability of a type I error and power test (0.05 significant level) of statistics on location testing between two populations with uniform distribution classified by sample size and variance

Sample Size	Var.ratio	t				Welch				t-BSE				WBR				B-M				Y-W				Wilcoxon	
		TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT	TE	PT		
10:10	1:1	<b>0.0506</b>	0.4267	<b>0.0494</b>	0.4241	<b>0.0544</b>	0.4379	<b>0.0521</b>	0.4009	<b>0.0547</b>	0.4246	<b>0.0541</b>	0.2523	<b>0.0429</b>	0.3668												
	1:2.25	<b>0.0546</b>	0.5620	<b>0.0529</b>	0.5516	<b>0.0573</b>	<b>0.5740</b>	<b>0.0590</b>	0.5014	<b>0.0570</b>	0.5105	<b>0.0592</b>	0.3355	<b>0.0496</b>	0.4674												
	1:4	<b>0.0595</b>	0.6312	<b>0.0548</b>	0.6072	<b>0.0631</b>	<b>0.6416</b>	<b>0.0678</b>	0.5213	<b>0.0564</b>	0.5050	<b>0.0676</b>	0.3724	<b>0.0572</b>	0.4882												
	1:9	<b>0.0624</b>	0.6804	<b>0.0547</b>	0.6386	<b>0.0655</b>	<b>0.6899</b>	<b>0.0746</b>	0.4949	<b>0.0524</b>	0.4331	0.0769	0.3798	<b>0.0690</b>	0.4743												
10:30	1:1	<b>0.0490</b>	0.6206	<b>0.0531</b>	0.5763	<b>0.0496</b>	<b>0.6248</b>	<b>0.0534</b>	0.5434	<b>0.0560</b>	0.5634	<b>0.0684</b>	0.3527	<b>0.0496</b>	0.5541												
	1:2.25	0.0190	0.7385	<b>0.0513</b>	<b>0.8463</b>	0.0204	0.7419	<b>0.0560</b>	0.7634	<b>0.0501</b>	0.7727	<b>0.0587</b>	0.5847	<b>0.0257</b>	0.6919												
	1:4	0.0088	0.7871	<b>0.0477</b>	<b>0.9406</b>	0.0089	0.7908	<b>0.0634</b>	0.8428	<b>0.0485</b>	0.8384	<b>0.0530</b>	0.7309	0.0179	0.7247												
	1:9	0.0043	0.8401	<b>0.0471</b>	<b>0.9846</b>	0.0045	0.8435	<b>0.0597</b>	0.9003	<b>0.0495</b>	0.9843	<b>0.0515</b>	0.8543	0.0158	0.7303												
30:30	1:1	<b>0.0500</b>	0.9132	<b>0.0498</b>																							



From table 9, it could be found that when the populations have uniform distribution and equal variance at 0.05 significant level, all test statistics could control type I errors in all sample sizes. For unequal variance cases, Welch t test and B-Mtest could control type I errors in all sample sizes followed by WBR test and Y-Wtest respectively.

It could be found that when a population had uniform distribution and equal variance, the test statistics with the highest testing power that could control type I errors were t-BSE test (85.71%). For unequal variance, the test statistics with the highest testing power that could control type I errors was Welch t test (71.43%).

**Table 10** The probability of a type I error and power test (0.01 significant level) of statistics on location testing between two populations with uniform distribution classified by sample size and variance

Sample Size	Var.ratio	t		Welch		t-BSE		WBR		B-M		Y-W		Wilcoxon	
		TE	PT	TE	PT	TE	PT								
10:10	1:1	<b>0.0124</b>	<b>0.1853</b>	<b>0.0116</b>	0.1801	<b>0.0099</b>	0.1595	<b>0.0118</b>	0.1769	0.0170	0.2430	<b>0.0148</b>	0.0838	<b>0.0093</b>	0.1511
	1:2.25	<b>0.0143</b>	<b>0.2890</b>	<b>0.0130</b>	0.2712	<b>0.0117</b>	0.2561	<b>0.0140</b>	0.2554	0.0176	0.3128	0.0152	0.1323	<b>0.0118</b>	0.2260
	1:4	0.0160	0.3432	<b>0.0138</b>	0.3028	<b>0.0135</b>	<b>0.3059</b>	0.0185	0.2752	0.0199	0.3001	0.0192	0.1523	<b>0.0150</b>	0.2463
	1:9	0.0205	0.3858	0.0166	0.3147	0.0170	0.3478	0.0220	0.2727	0.0200	0.2565	0.0236	0.1654	0.0198	0.2550
10:30	1:1	<b>0.0095</b>	<b>0.3529</b>	<b>0.0135</b>	0.2900	<b>0.0085</b>	0.3390	<b>0.0118</b>	0.2931	<b>0.0150</b>	0.3324	0.0241	0.1574	<b>0.0088</b>	0.2922
	1:2.25	0.0032	0.4184	<b>0.0140</b>	<b>0.6042</b>	0.0023	0.3998	0.0158	0.5159	0.0152	0.5650	0.0165	0.3090	0.0037	0.3871
	1:4	0.0010	0.4500	<b>0.0099</b>	<b>0.7917</b>	0.0010	0.4285	0.0156	0.6363	<b>0.0126</b>	0.6553	<b>0.0123</b>	0.4456	0.0026	0.4087
	1:9	0.0002	0.4728	<b>0.0109</b>	<b>0.9228</b>	0.0001	0.4514	0.0228	0.7367	<b>0.0110</b>	0.6882	<b>0.0131</b>	0.6138	0.0019	0.4076
30:10	1:1	<b>0.0096</b>	<b>0.3659</b>	<b>0.0130</b>	0.2935	<b>0.0094</b>	0.3473	<b>0.0132</b>	0.2939	0.0158	0.3363	0.0180	0.1591	<b>0.0087</b>	0.2953
	1:2.25	0.0362	0.6335	<b>0.0146</b>	<b>0.3224</b>	0.0339	0.6205	0.0156	0.3368	0.0181	0.3537	0.0265	0.1756	0.0252	0.4226
	1:4	0.0622	0.7469	<b>0.0134</b>	<b>0.3295</b>	0.0587	0.7329	0.0156	0.3262	0.0153	0.3222	0.0218	0.1695	0.0345	0.4523
	1:9	0.1086	0.8266	0.0171	0.3240	0.1059	0.8171	0.0230	0.2919	0.0204	0.2688	0.0257	0.1621	0.0563	0.4618
30:30	1:1	<b>0.0096</b>	<b>0.7514</b>	<b>0.0096</b>	0.7512	<b>0.0091</b>	0.7413	<b>0.0098</b>	0.6779	<b>0.0117</b>	0.7090	<b>0.0101</b>	0.4263	<b>0.0088</b>	0.6665
	1:2.25	<b>0.0106</b>	<b>0.9009</b>	<b>0.0106</b>	0.8979	<b>0.0100</b>	0.8968	<b>0.0115</b>	0.8042	<b>0.0103</b>	0.8168	<b>0.0120</b>	0.5951	<b>0.0106</b>	0.7971
	1:4	<b>0.0096</b>	<b>0.9488</b>	<b>0.0091</b>	0.9443	<b>0.0090</b>	0.9454	<b>0.0135</b>	0.8257	<b>0.0104</b>	0.8139	<b>0.0110</b>	0.6699	<b>0.0130</b>	0.8191
	1:9	<b>0.0104</b>	<b>0.9681</b>	<b>0.0087</b>	0.9629	<b>0.0096</b>	0.9651	0.0163	0.8028	<b>0.0091</b>	0.7552	<b>0.0127</b>	0.7045	0.0159	0.8000
50:100	1:1	<b>0.0109</b>	<b>0.9912</b>	<b>0.0116</b>	0.9909	<b>0.0113</b>	<b>0.9912</b>	<b>0.0110</b>	0.9762	<b>0.0117</b>	0.9785	<b>0.0111</b>	0.8622	<b>0.0110</b>	0.9742
	1:2.25	0.0032	0.9997	<b>0.0092</b>	<b>1.0000</b>	0.0034	0.9996	<b>0.0122</b>	0.9986	<b>0.0104</b>	0.9986	<b>0.0104</b>	0.9813	0.0048	0.9970
	1:4	0.0016	0.9998	<b>0.0095</b>	<b>1.0000</b>	0.0018	0.9998	0.0154	0.9995	<b>0.0094</b>	0.9991	<b>0.0107</b>	0.9958	0.0036	0.9981
	1:9	0.0007	1.0000	<b>0.0103</b>	<b>1.0000</b>	0.0007	1.0000	0.0236	0.9999	<b>0.0108</b>	0.9995	<b>0.0116</b>	0.9996	<b>0.0051</b>	0.9984
100:50	1:1	<b>0.0095</b>	0.9930	<b>0.0098</b>	<b>0.9932</b>	<b>0.0092</b>	0.9928	<b>0.0100</b>	0.9808	<b>0.0111</b>	0.9823	<b>0.0117</b>	0.8624	<b>0.0092</b>	0.9788
	1:2.25	0.0231	0.9993	<b>0.0098</b>	<b>0.9982</b>	0.0221	0.9991	<b>0.0116</b>	0.9899	<b>0.0102</b>	0.9902	<b>0.0107</b>	0.9249	0.0196	0.9932
	1:4	0.0383	0.9999	<b>0.0099</b>	<b>0.9992</b>	0.0375	0.9998	<b>0.0148</b>	0.9809	<b>0.0113</b>	0.9786	<b>0.0141</b>	0.9434	0.0309	0.9895
	1:9	0.0551	1.0000	<b>0.0100</b>	<b>1.0000</b>	0.0536	1.0000	0.0171	0.9705	<b>0.0110</b>	0.9614	<b>0.0118</b>	0.9609	0.0419	0.9886
100:100	1:1	<b>0.0095</b>	<b>0.9996</b>	<b>0.0095</b>	0.9996	<b>0.0088</b>	<b>0.9996</b>	<b>0.0095</b>	0.9979	<b>0.0100</b>	0.9980	<b>0.0095</b>	0.9726	<b>0.0093</b>	0.9978
	1:2.25	<b>0.0101</b>	<b>1.0000</b>	<b>0.0099</b>	<b>1.0000</b>	<b>0.0093</b>	<b>1.0000</b>	<b>0.0118</b>	<b>1.0000</b>	<b>0.0106</b>	<b>1.0000</b>	<b>0.0109</b>	0.9977	<b>0.0115</b>	<b>1.0000</b>
	1:4	<b>0.0099</b>	<b>1.0000</b>	<b>0.0098</b>	<b>1.0000</b>	<b>0.0094</b>	<b>1.0000</b>	<b>0.0142</b>	<b>1.0000</b>	<b>0.0101</b>	<b>1.0000</b>	<b>0.0105</b>	0.9995	<b>0.0139</b>	<b>1.0000</b>
	1:9	<b>0.0098</b>	<b>1.0000</b>	<b>0.0094</b>	<b>1.0000</b>	<b>0.0099</b>	<b>1.0000</b>	0.0183	0.9994	<b>0.0102</b>	0.9994	<b>0.0112</b>	0.9995	0.0181	0.9994

\* Italic and bold indicate the case that was able to control type I error.

\*\* Bold entries indicate the highest testing power controlled under a type I error.

From table 10, it could be found that when the populations have uniform distribution and equal variance at 0.01 significant level, all test statistics can control type I errors in all sample sizes except B-Mtest and Y-W test. For unequal variance cases Welch t test can control type I errors in almost sample sizes followed by BM test and YWtest.

It could be found that when the population has uniform distribution and equal variance, the test statistics with the highest testing power that could control type I errors is t test (85.71%). For unequal variance, the test statistics with the highest testing power that could control some sample sizes was Welch t test (66.67%).

## Discussion

The results revealed that when the populations had five distributions on equal variance, all test statistics could control the type I error on almost sample sizes. This means that the said test statistics were robust. Nonetheless, when the variance was unequal, it was found that all test statistics had a lower capacity in controlling the type I error. Specifically, when the populations had log-normal distribution, gamma



distribution and exponential distribution on unequal variance ratio was 1:9, it could control the type I error on some sample sizes. Moreover, the results also showed that the test statistics of WBR test, B-M test, Y-W test, and Exact Wilcoxon signed-rank test could not control the type I error when the variance was unequal in almost all conditions. The above test statistics were proposed to solve the problem of Behrenses-Fisher. As shown in table 4, the ratio of variances was 1:9 with the sample size of 100,100 and the type I error was found at above 23.35%. Comparing with Welch t test, the type I error was found at approximately 1.56%. In particular, the Exact Wilcoxon signed-rank test statistic is not suitable for hypotheses testing when the variance is unequal (Harwell et al., 1992; Zimmerman & Zumbo, 1993a; Zimmerman & Zumbo, 1993b; Stonehouse & Forrester, 1998). Furthermore, the test statistics with the highest testing power that could control the type I error for equal variance was WBR while for unequal variance was Welch t test. This maybe caused by the said location testing was the test of the difference of the mean which, the Welch t test was therefore more efficient than other test statistics.

### **Conclusion and Suggestions**

In summary , it could be said that when the populations had five distributions on equal variance, all test statistics could control the type I error on almost sample sizes, both equal and unequal sizes. The test statistics with the highest testing power that could control type I errors was WBR test. In the case of unequal variance with the WBR test, B-M test, Y-W test, and Exact Wilcoxon signed-rank test could not control the type I error in almost all conditions especially when the population was enumerated log-normal distribution, gamma distribution and exponential distribution. However, when considering together both the capacity to control the type I error and highest testing power, the Welch was test is the best test statistic. Therefore, the application of the said test statistics must be carefully done in real cases.

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